

NORMAL APPROXIMATION OF U -STATISTICS IN HILBERT SPACE*

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Abstract. Let $\{U_n\}$, $n = 1, 2, \dots$, be Hilbert space H -valued U -statistics with kernel $\Phi(\cdot, \cdot)$, corresponding to a sequence of observations (random variables) X_1, X_2, \dots . The rate of convergence on balls in the central limit theorem for $\{U_n\}$ is investigated. The obtained estimate is of order $n^{-1/2}$ and depends explicitly on $\mathbf{E}\|\Phi(X_1, X_2)\|^3$ and on the trace and the first nine eigenvalues of the covariance operator of $\mathbf{E}(\Phi(X_1, X_2)|X_1)$.

Key words. U -statistic, Hilbert space, central limit theorem, normal (Gaussian) approximation, rate of convergence

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1. Introduction. Let X_1, \dots, X_n be independent random variables with values in a measurable space (X, \mathcal{X}) having on it the same distribution \mathbf{P} . Let H be a real separable Hilbert space with inner product (\cdot, \cdot) and norm $\|\cdot\|$. Consider the UH -statistic

$$U_n = \binom{n}{2}^{-1} \sum_{1 \leq i < j \leq n} \Phi(X_i, X_j)$$

with a symmetric kernel $\Phi: X^2 \rightarrow H$ such that $\mathbf{E}\Phi(X_1, X_2) = 0$, $\mathbf{E}\|\Phi(X_1, X_2)\| < \infty$.

By the Hoeffding decomposition,

$$(1) \quad U_n = 2n^{-1} \sum_{j=1}^n g_1(X_j) + 2n^{-1}(n-1)^{-1} \sum_{1 \leq i < j \leq n} g_2(X_i, X_j),$$

where

$$g_1(x) = \int_X \Phi(x, y) \mathbf{P}(dy) = \mathbf{E}(\Phi(X_1, X_2) | X_1 = x),$$

$$g_2(x, y) = \Phi(x, y) - g_1(x) - g_1(y).$$

We assume that $\sigma^2 = \mathbf{E}\|g_1(X_1)\|^2 > 0$ and observe that the definitions of $g_1(x)$ and $g_2(x)$ imply

$$(2) \quad \int_X g_2(x, y) \mathbf{P}(dy) = 0 \quad (\text{a.s.}).$$

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Assuming also that $\sigma < \infty$ we can define the covariance operator V of $g_1(X_1)$ in the usual way:

$$(Vx, y) = \int_X (x, g_1(z)) (y, g_1(z)) \mathbf{P}(dz), \quad x, y \in H.$$

Let Y be a Gaussian H -valued random variable with mean zero and covariance operator $\sigma^{-2}V$. Denote by $\sigma_1^2 \geq \sigma_2^2 \geq \dots$ the eigenvalues of V . Finally, for any $r \geq 0$, $a \in H$, let $B_r(a) = \{x \in H: \|x - a\| < r\}$ and $B_r = B_r(0)$.

By the central limit theorem for UH -statistics, if

$$\mathbf{E}\|g_1(X_1)\|^2 < \infty, \quad \mathbf{E}\|g_2(X_1, X_2)\|^{4/3} < \infty,$$

then (see [1], [2], [5])

$$\Delta_n(a) = \sup_r \left| \mathbf{P}\{\|2^{-1}n^{1/2}\sigma^{-1}U_n - a\| < r\} - \mathbf{P}\{\|Y - a\| < r\} \right| \longrightarrow 0, \quad n \rightarrow \infty.$$

The rate of convergence of $\Delta_n(0)$ to zero with respect to n was estimated in [1], [2], [5], [12], where under the additional moment conditions

$$\mathbf{E}\|g_1(X_1)\|^3 < \infty, \quad \mathbf{E}\|g_2(X_1, X_2)\|^2 < \infty$$

it was shown that

$$\Delta_n(0) = \rho_n(V) n^{-1/2}$$

with $\rho_n(V) \rightarrow 0$ as $n \rightarrow \infty$. To prove this, it was observed that the problem of estimating the probability for a value of the UH -statistic to be in the ball B_r may be treated as the problem of estimating the rate of convergence in the central limit theorem in Hilbert space for balls with a nonzero random center which tends to zero with probability 1 as $n \rightarrow \infty$. However, it was assumed in [1], [2], [5], [12] that the number of nonzero eigenvalues of the covariance operator V is sufficiently large. The explicit dependence of $\rho_n(V)$ on V in [1], [2], [5], [12] was not determined. Estimates of order $O(n^{-1/3})$ for any $a \in H$ with explicit dependence on the second and third moments were obtained in [14].

Note that in the one-dimensional case the study of the problem considered in the present paper has a rather long history. It was initiated in 1973 by Grams and Serfling [11] and later a number of gradual improvements of their result was made (for details see [13] and [10]). In [10] it was also shown that the estimate obtained by Korolyuk and Borovskikh in [4] is essentially unimprovable.

In the present paper we study the dependence of $\Delta_n(a)$ on n , the covariance operator V and $\beta = \mathbf{E}\|\Phi(X_1, X_2)\|^3$. When $\beta < \infty$ an estimate of order $n^{-1/2}$ is obtained, the dependence of which on Φ involves only β , the first nine eigenvalues of V and $\sigma^2 = \text{tr } V$. The order of dependence of our estimate on n is precise, but its dependence on σ_j^2, σ_j^2 and β apparently admits improvement. To prove the estimate we employ the methods developed in [3] and [12].

2. The theorem.

THEOREM. *There exists an absolute constant c such that, for all $n \geq 2$ and $a \in H$,*

$$(3) \quad \Delta_n(a) \leq c \left(1 + \|a\|^3\right) \sigma^4 \left(\prod_{j=1}^9 \sigma_j^{-1}\right) \sigma_9^{-1} \beta^2 n^{-1/2}.$$

In what follows we will assume that $\sigma = 1$. The general case is reduced to this one if we replace Φ by $\sigma^{-1}\Phi$. We can assume also that $\beta < \infty$ and $\sigma_g > 0$ since otherwise (3) would be obvious.

In the proof of the theorem the upper estimates for $\Delta_n(a)$ are calculated more precisely; except for n and V , they depend only on the moments

$$\mathbf{E}\|g_1\|^3 = \mathbf{E}\|g_1(X_1)\|^3, \quad \mathbf{E}\|g_2\|^3 = \mathbf{E}\|g_2(X_1, X_2)\|^3,$$

which are related to β by the inequalities

$$(4) \quad \mathbf{E}\|g_1\|^3 \leq \beta, \quad \mathbf{E}\|g_2\|^3 \leq 27\beta.$$

We formulate the theorem using only β for the sake of simplicity.

3. Proof of the theorem. Below c, c_1, c_2, \dots denote absolute constants; the same symbol may stand for different constants. The proof is divided into several lemmas.

LEMMA 1. *We have*

$$(5) \quad \Delta_n(a) \leq \Delta_{n1}(a) + n\mathbf{P}\{\|g_1(X_1)\| > n^{1/2}\},$$

where

$$\Delta_{n1}(a) = \sup_r \left| \mathbf{P}\{\|\bar{S}_n - a + V_n\| < r\} - \mathbf{P}\{\|Y - a\| < r\} \right|,$$

$$\bar{S}_n = n^{-1/2} \sum_{j=1}^n g_1(X_j) I\left(\|g_1(X_j)\| \leq n^{1/2}\right),$$

$I(A)$ is the indicator function of the set A and

$$V_n = (n-1)^{-1} n^{-1/2} \sum_{j=1}^n \sum_{i=1}^{j-1} g_2(X_i, X_j).$$

Proof. According to (1),

$$2^{-1} n^{1/2} U_n - a = S_n - a + V_n,$$

where

$$S_n = n^{-1/2} \sum_{j=1}^n g_1(X_j)$$

and we obviously have

$$\Delta_n(a) \leq \Delta_{n1}(a) + \sup_r \left| \mathbf{P}\{\|S_n - a + V_n\| < r\} - \mathbf{P}\{\|\bar{S}_n - a + V_n\| < r\} \right|.$$

Now applying Lemma 1 of [17] with $f(x) = I(\|x\| < r)$, $u_1 = \infty$, $u_2 = n^{1/2}$ and $Z = -a + V_n$, we see that the second summand in the right-hand side of this inequality is not greater than $n\mathbf{P}\{\|g_1(X_1)\| > n^{1/2}\}$ and the lemma follows.

Let $k = n - [n^{2/3}]$, where $[n^{2/3}]$ is the integral part of $n^{2/3}$ and denote

$$\begin{aligned}\alpha_k &= \sum_{1 \leq i < j \leq k} g_2(X_i, X_j), & \beta_k &= \sum_{i=1}^k \sum_{j=k+1}^n g_2(X_i, X_j), \\ \delta_k &= \sum_{k+1 \leq i < j \leq n} g_2(X_i, X_j), & \bar{S}_n^k &= n^{-1/2} \sum_{j=1}^k \bar{g}_1(X_j), \\ \bar{g}_1(X_j) &= g_1(X_j) I\left(\|g_1(X_j)\| \leq n^{1/2}\right), & j &= 1, \dots, n, \\ Z_n &= 2\varepsilon_n(\bar{S}_n^k - a, \alpha_k), & \varepsilon_n &= (n-1)^{-1} n^{-1/2}.\end{aligned}$$

LEMMA 2. For $\Delta_{n1}(a)$ we have

$$(6) \quad \Delta_{n1}(a) \leq \Delta_{n2}(a) + c_1(1 + \|a\|^3) (\beta^2 + \sigma_1^{-1} \sigma_2^{-1}) n^{-1/2},$$

where

$$\Delta_{n2}(a) = \sup_r \left| \mathbf{P}\{\|\bar{S}_n - a\|^2 + Z_n < r\} - \mathbf{P}\{\|Y - a\|^2 < r\} \right|.$$

Proof. First we observe that

$$(7) \quad \begin{aligned}\|\bar{S}_n - a + V_n\|^2 &= (\bar{S}_n - a + V_n, \bar{S}_n - a + V_n) \\ &= \|\bar{S}_n - a\|^2 + 2(\bar{S}_n - a, V_n) + \|V_n\|^2.\end{aligned}$$

In our notation,

$$\begin{aligned}\|V_n\|^2 &= (n-1)^{-2} n^{-1} \left\| \sum_{j=2}^n \sum_{i=1}^{j-1} g_2(X_i, X_j) \right\|^2, \\ 2(\bar{S}_n - a, V_n) &= Z_n + 2\varepsilon_n(\bar{S}_n^k - a, \beta_k + \delta_k) + 2(\bar{S}_n - \bar{S}_n^k, V_n),\end{aligned}$$

and

$$\begin{aligned}(\bar{S}_n^k - a, \beta_k + \delta_k) &= (\bar{S}_n^k, \beta_k) - (a, \beta_k) + (\bar{S}_n^k, \delta_k) - (a, \delta_k), \\ (\bar{S}_n - \bar{S}_n^k, V_n) &= \varepsilon_n(\bar{S}_n - \bar{S}_n^k, \alpha_k) + \varepsilon_n(\bar{S}_n - \bar{S}_n^k, \beta_k) + \varepsilon_n(\bar{S}_n - \bar{S}_n^k, \delta_k).\end{aligned}$$

In what follows we will need the following inequalities: for any $p \geq 2$,

$$(8) \quad \mathbf{E}\|\bar{S}_n^k\|^p \leq c_p, \quad \left\| \mathbf{E} \bar{g}_1(X_1) \right\| = \left\| \int_{\|g_1(x)\| > n^{1/2}} g_1(x) \mathbf{P}(dx) \right\| \leq n^{-1/2}$$

(see [9], [16]) and

$$(9) \quad \begin{aligned}\mathbf{E}\|\alpha_k\|^p &\leq c_p \mathbf{E}\|g_2\|^p k^p, \\ \mathbf{E}\|\beta_k\|^p &\leq c_p \mathbf{E}\|g_2\|^p k^{p/2} (n-k)^{p/2}, \\ \mathbf{E}\|\delta_k\|^p &\leq c_p \mathbf{E}\|g_2\|^p (n-k)^p\end{aligned}$$

(see [5]).

The Chebychev inequality and (9) imply

$$\mathbf{P}\{\|V_n\|^2 \geq n^{-1/2}\} \leq n^{1/2} \mathbf{E}\|V_n\|^2 \leq c_1 \mathbf{E}\|g_2\|^2 n^{-1/2}.$$

Furthermore,

$$(10) \quad \mathbf{P}\left\{|\overline{S}_n^k, \varepsilon_n \beta_k| > n^{-1/2}\right\} \leq \Pi_1 + \Pi_2,$$

where

$$\begin{aligned} \Pi_1 &= \mathbf{P}\left\{2k\varepsilon_n \left| \mathbf{E}\overline{g}_1(X_1), \sum_{i=1}^k h(X_i) \right| > 1\right\}, \\ \Pi_2 &= \mathbf{P}\left\{2\varepsilon_n \left| \sum_{i=1}^k \overline{g}(X_i), \sum_{i=1}^k h(X_i) \right| > 1\right\}, \end{aligned}$$

$$\overline{g}(X_i) = \overline{g}_1(X_i) - \mathbf{E}\overline{g}_1(X_1), \quad h(X_i) = \sum_{j=k+1}^n g_2(X_i, X_j), \quad i = 1, \dots, k.$$

For Π_1 by the Chebychev inequality and (8), (9)

$$\Pi_1 \leq (2k\varepsilon_n)^3 \|\mathbf{E}g_1(X_1)\|^3 \mathbf{E}\left\|\sum_{i=1}^k h(X_i)\right\|^3 \leq c_1 \mathbf{E}\|g_2\|^3 n^{-1/2}.$$

To estimate Π_2 we first observe that

$$\sum_{i=1}^k \overline{g}(X_i), \sum_{i=1}^k h(X_i) = \sum_{i=1}^4 A_i,$$

where

$$\begin{aligned} A_1 &= k \mathbf{E}_1(\overline{g}(X_1), h(X_1)), \\ A_2 &= \sum_{i=1}^k \left((\overline{g}(X_i), h(X_i)) - \mathbf{E}_1(\overline{g}(X_1), h(X_1)) \right), \\ A_3 &= \sum_{i_1=2}^k \sum_{i_2=1}^{i_1-1} (\overline{g}(X_{i_1}), h(X_{i_2})), \quad A_4 = \sum_{i_1=1}^{k-1} \sum_{i_2=i_1+1}^k (\overline{g}(X_{i_1}), h(X_{i_2})) \end{aligned}$$

and, obviously,

$$\mathbf{E}_k |A_1|^3 \leq k^3 \left| \mathbf{E}_1(\overline{g}(X_1), h(X_1)) \right|^3$$

by [8]

$$\mathbf{E}_k |A_2|^3 \leq c_1 k^{3/2} \mathbf{E}_1 \left| (\overline{g}(X_1), h(X_1)) \right|^3$$

and by [9]

$$\mathbf{E}_k |A_3|^3 \leq c_1 k^3 \mathbf{E}_2 \left| (\overline{g}(X_1), h(X_2)) \right|^3, \quad \mathbf{E}_k |A_4|^3 \leq c_1 k^3 \mathbf{E}_2 \left| (\overline{g}(X_1), h(X_2)) \right|^3,$$

$\mathbf{E}_k(\cdot)$ denotes the expectation with respect to X_1, \dots, X_k with fixed X_{k+1}, \dots, X_n , $k = 1, \dots, n$.

Moreover, since $\mathbf{E}\|g_1(X_1)\|^2 = 1$, we have

$$\left| \mathbf{E}(\overline{g}(X_1), h(X_1)) \right|^3 \leq \mathbf{E}_1 \|h(X_1)\|^3.$$

And, obviously,

$$\begin{aligned}\mathbf{E}_1 \left| (\bar{g}(X_1), h(X_1)) \right|^3 &\leq 8n^{3/2} \mathbf{E}_1 \|h(X_1)\|^3, \\ \mathbf{E}_2 \left| (\bar{g}(X_1), h(X_2)) \right|^3 &\leq c_1 \mathbf{E}_1 \|\bar{g}(X_1)\|^3 \mathbf{E}_1 \|h(X_1)\|^3\end{aligned}$$

and by [8]

$$\bar{\mathbf{E}}_k \|h(X_1)\|^3 \leq c_1 (n-k)^{3/2} \bar{\mathbf{E}}_k \|g_2(X_1, X_{k+1})\|^3,$$

$\bar{\mathbf{E}}_k(\cdot)$ denotes the expectation with respect to X_{k+1}, \dots, X_n with fixed X_1, \dots, X_k , $k = 1, \dots, n$.

The above implies

$$\Pi_2 \leq c_1 \mathbf{E} \|g_1\|^3 \mathbf{E} \|g_2\|^3 n^{-1/2}$$

and, hence we obtain an estimate for the probability in the left side of (10).

Similarly, using (8), (9) we get

$$\begin{aligned}\mathbf{P} \left\{ |c(\bar{S}_n^k, \varepsilon_n \delta_k)| > n^{-1/2} \right\} &\leq \mathbf{P} \left\{ c \|\bar{S}_n^k\| > n^{1/6} \right\} + \mathbf{P} \left\{ n^{1/6} \varepsilon_n \|\delta_k\| > n^{-1/2} \right\} \\ &\leq c^3 \mathbf{E} \|\bar{S}_n^k\|^3 n^{-1/2} + n^{1/2} n^{3/2} \varepsilon_n^3 \mathbf{E} \|\delta_k\|^3 \leq c_1 n^{-1/2} + c_2 \mathbf{E} \|g_2\|^3 n^{-1/2}, \\ \mathbf{P} \left\{ |c(a, \varepsilon_n \beta_k)| > n^{-1/2} \right\} &\leq c^3 \|a\|^3 \varepsilon_n^3 n^{3/2} \mathbf{E} \|\beta_k\|^3 \leq c_1 \mathbf{E} \|a\|^3 \mathbf{E} \|g_2\|^3 n^{-1/2}, \\ \mathbf{P} \left\{ |c(a, \varepsilon_n \delta_k)| > n^{-1/2} \right\} &\leq c^3 \|a\|^3 \varepsilon_n^3 n^{3/2} \mathbf{E} \|\delta_k\|^3 \leq c_1 \mathbf{E} \|a\|^3 \mathbf{E} \|g_2\|^3 n^{-1/2}, \\ \mathbf{P} \left\{ |c(\bar{S}_n - \bar{S}_n^k, \varepsilon_n \alpha_k)| > n^{-1/2} \right\} &\leq c^3 n^{3/2} \mathbf{E} \|\bar{S}_n - \bar{S}_n^k\|^3 \varepsilon_n^3 \mathbf{E} \|\alpha_k\|^3 \\ &\leq c_1 \mathbf{E} \|g_1\|^3 \mathbf{E} \|g_2\|^3 n^{-1/2}, \\ \mathbf{P} \left\{ |c(\bar{S}_n - \bar{S}_n^k, \varepsilon_n \beta_k)| > n^{-1/2} \right\} &\leq \mathbf{P} \left\{ c \|\bar{S}_n - \bar{S}_n^k\| > 1 \right\} + \mathbf{P} \left\{ \varepsilon_n \|\beta_k\| > n^{-1/2} \right\} \\ &\leq c^3 \mathbf{E} \|\bar{S}_n - \bar{S}_n^k\|^3 + n^{3/2} \varepsilon_n^3 \mathbf{E} \|\beta_k\|^3 \leq c_1 \mathbf{E} \|g_1\|^3 n^{-1/2} + c_2 \mathbf{E} \|g_2\|^3 n^{-1/2}, \\ \mathbf{P} \left\{ |c(\bar{S}_n - \bar{S}_n^k, \varepsilon_n \delta_k)| > n^{-1/2} \right\} &\leq \mathbf{P} \left\{ c \|\bar{S}_n - \bar{S}_n^k\| > 1 \right\} + \mathbf{P} \left\{ \varepsilon_n \|\delta_k\| > n^{-1/2} \right\} \\ &\leq c^3 \mathbf{E} \|\bar{S}_n - \bar{S}_n^k\|^3 + n^{3/2} \varepsilon_n^3 \mathbf{E} \|\delta_k\|^3 \leq c_1 \mathbf{E} \|g_1\|^3 n^{-1/2} + c_2 \mathbf{E} \|g_2\|^3 n^{-1/2}.\end{aligned}$$

Denote now

$$Y_n = 2\varepsilon_n (\bar{S}_n^k - a, \beta_k + \delta_k) + 2(\bar{S}_n - \bar{S}_n^k, V_n) + \|V_n\|^2.$$

Then, according to (7),

$$\|\bar{S}_n - a + V_n\|^2 = \|\bar{S}_n - a\|^2 + Z_n + Y_n$$

and the estimates above imply

$$(11) \quad \mathbf{P} \left\{ |Y_n| > n^{-1/2} \right\} \leq c_1 (1 + \|a\|^3) \mathbf{E} \|g_1\|^3 \mathbf{E} \|g_2\|^3 n^{-1/2}.$$

Since, for all $r \geq 0$,

$$\begin{aligned}\mathbf{P} \left\{ \|\bar{S}_n - a\|^2 + Z_n < r - n^{-1/2} \right\} &- \mathbf{P} \left\{ |Y_n| > n^{-1/2} \right\} \\ &\leq \mathbf{P} \left\{ \|\bar{S}_n - a\|^2 + Z_n + Y_n < r \right\} \\ &\leq \mathbf{P} \left\{ \|\bar{S}_n - a\|^2 + Z_n < r + n^{-1/2} \right\} + \mathbf{P} \left\{ |Y_n| > n^{-1/2} \right\},\end{aligned}$$

we have

$$(12) \quad \Delta_{n1}(a) \leq \Delta_{n2}(a) + \mathbf{P}\{|Y_n| > n^{-1/2}\} + \sup_r \mathbf{P}\{r - n^{-1/2} \leq \|Y - a\|^2 \leq r + n^{-1/2}\}.$$

In [6] it is proved that

$$(13) \quad \sup_r \mathbf{P}\{r - n^{-1/2} \leq \|Y - a\|^2 \leq r + n^{-1/2}\} \leq c\sigma_1^{-1}\sigma_2^{-1}n^{-1/2}.$$

Estimates (11)–(13) imply (6).

Denote now

$$\begin{aligned} \varphi_n(t) &= \mathbf{E} \exp(it\|\bar{S}_n - a\|^2 + itZ_n), \\ \varphi(t) &= \mathbf{E} \exp(it\|Y - a\|^2), \\ \psi_n(t) &= \mathbf{E} \exp(it\|\bar{S}_n - a\|^2). \end{aligned}$$

The density of the random variable $\|Y - a\|^2$ is not greater than $c(\sigma_1\sigma_2)^{-1}$ (see [6]). Thus for $\Delta_{n2}(a)$ in (6), applying Esséen’s inequality (see, e.g. [7]) for any $T > 0$ we have

$$(14) \quad \Delta_{n2}(a) \leq c_1 \int_{-T}^T |\varphi_n(t) - \varphi(t)| |t|^{-1} dt + c_2(\sigma_1\sigma_2)^{-1}T^{-1}.$$

We choose in (14)

$$(15) \quad T = \frac{cn^{1/2}}{L}, \quad L = 3\sigma_9^{-2}\mathbf{E}\|g_1\|^3, \quad \sigma_9 > 0,$$

where c is a small enough absolute constant and then apply Lemma 12 of [3]. Note that with L as in (15) the following conditions of Lemma 12 in [3]

$$(16) \quad \frac{2}{3} \leq L^2 \leq \frac{n}{2e}, \quad \mathbf{P}(B_L) \geq \frac{1}{2}$$

are fulfilled (for a proof see [3], [9], [16]). When (16) is satisfied the distribution of $\bar{g}_1 = g_1(X_1)I(\|g_1(X_1)\| \leq n^{1/2})$ may be represented as $(Q + Q_1)/2$, where

$$Q(A) = \frac{\mathbf{P}(A \cap B_L)}{\mathbf{P}(B_L)}, \quad A \in H,$$

and Q_1 is a probability measure (see [15]).

Let V_1 be the covariance operator of Q and $\sigma_{11}^2 \geq \sigma_{12}^2 \geq \dots$ be its eigenvalues. By our assumptions, $\text{tr } V = 1$. Hence (see [3]) $\text{tr } V_1 \leq 2$ and

$$(17) \quad \frac{2\sigma_j^2}{3} \leq \sigma_{1j}^2 \leq 2\sigma_j^2, \quad j = 1, \dots, 9.$$

To estimate the integral in (14) we will need the following generalization of Lemma 11 in [3] (see also [9, Lemma 4]).

LEMMA 3. *Denote*

$$Z = n^{-1/2} \sum_{j=1}^{l+m} \bar{g}_1(X_j), \quad Z_1(f) = \sum_{j=1}^{l+m} f(X_j),$$

where $\bar{g}_1 = g_1(X_1) I(\|g_1(X_1)\| \leq n^{1/2})$ and $f: X \rightarrow \mathbf{R}$ is a measurable function, l, m are positive integers, $l \leq m, l + m \leq n$. Then, for any $A > 0$, even $k \geq 0$, integer $r \geq 0, k_j \geq 0, j = 1, \dots, r$, and any $x, x_j \in H, j = 1, \dots, r$, for $|t| \leq c(A) nL^{-1}(l \log(l/L^2))^{-1/2}$ if $l > L^2$, and for any t if $l \leq L^2$

$$(18) \quad \left| \mathbf{E} \exp(it\|Z + x\|^2 + it Z_1(F)) \|Z\|^k \prod_{j=1}^r (x_j, Z)^{k_j} \right| \leq c_1 \prod_{j=1}^r \|x_j\|^{k_j} \exp(-c_2 l) + c_3 \left(\frac{L^2}{l}\right)^A + h^{1/2} \left(\frac{c_4 T_1^2 l m}{n^2}, V_1\right),$$

where c_1, c_2 are functions of $\bar{k} = k + k_1 + \dots + k_r$; c_3, c_4 are functions of A and \bar{k}

$$T_1 = \min \left\{ |t|, nL^{-1} \left(m \log \left(\frac{m}{L^2}\right)\right)^{-1/2} \right\}, \quad h(s, V_1) = \prod_{j=1}^{\infty} (1 + 2s\sigma_{1j}^4)^{-1/2}.$$

Proof. This lemma reduces to Lemma 11 in [3] if $Z_1(f) = 0$. The proof of the present lemma is essentially the same as the proof of Lemma 11 in [3].

Let us estimate the integral in (14). Obviously,

$$(19) \quad \int_{-T}^T |\varphi_n(t) - \varphi(t)| |t|^{-1} dt \leq I_1 + I_2,$$

where

$$I_1 = \int_{-T}^T | \varphi_n(t) - \varphi(t) | |t|^{-1} dt, \quad I_2 = \int_{-T}^T |\varphi_n(t) - \psi_n(t)| |t|^{-1} dt.$$

By Lemma 12 in [3]

$$(20) \quad I_1 \leq c_1 U,$$

where

$$U = c_7(V) \left(\mathbf{E}(\|g_1\|^2 + (a, g_1)^2) I(\|g_1\| > n^{1/2}) \right) + \left(\mathbf{E}(\|g_1\|^3 + |(a, g_1)|^3) I(\|g_1\| \leq n^{1/2}) + (Va, a)^{3/2} \right) n^{-1/2} + c_3(V) \sigma_9^{-2} \mathbf{E}\|g_1\|^3 n^{-1/2}, \quad c_s(V) = \prod_{j=1}^s \sigma_j^{-1},$$

and for U we have

$$U \leq (1 + \|a\|^3) c_7(V) \mathbf{E}\|g_1\|^3 n^{-1/2} + c_3(V) \sigma_9^{-2} \mathbf{E}\|g_1\|^3 n^{-1/2} \leq 2(1 + \|a\|^3) \beta c_9(V) \sigma_9^{-1} n^{-1/2}.$$

Consider now I_2 in (19). Obviously,

$$\varphi_n(t) - \psi_n(t) = \mathbf{E} \exp \{ it \|\bar{S}_n - a\|^2 \} (e^{itZ_n} - 1).$$

Next we will use the representation

$$(21) \quad e^z - 1 = \sum_{s=1}^{r-1} \frac{z^s}{s!} + z^r \int_0^1 \frac{(1-u)^{r-1} e^{uz}}{(r-1)!} du,$$

valid for all complex z and positive integers r , with $r = 3$, $z = it Z_n$. Thus we have

$$(22) \quad \varphi_n(t) - \psi_n(t) = it J_1(t) - \frac{t^2}{2} J_2(t) + \frac{(it)^3}{2} J_3(t),$$

where

$$J_1(t) = \mathbf{E}\left(Z_n \exp\{it \|\bar{S}_n - a\|^2\}\right), \quad J_2(t) = \mathbf{E}\left(Z_n^2 \exp\{it \|\bar{S}_n - a\|^2\}\right),$$

$$J_3(t) = \int_0^1 du (1-u)^2 \mathbf{E}\left(Z_n^3 \exp\{it \|\bar{S}_n - a\|^2 + itu Z_n\}\right).$$

LEMMA 4. *If $|t| \leq T = cn^{1/2}L^{-1}$, then*

$$(23) \quad |J_3(t)| \leq c_1 L^{3/2} n^{-1/2} + \prod_{j=1}^7 \left(1 + \frac{c_2 \sigma_j^4 t^2}{n^{2/3}}\right)^{-1/4}$$

$$\times (\mathbf{E} \|g_1\|^3 + \|a\|^3) \mathbf{E} \|g_2\|^3 n^{-3/2}.$$

Proof. The random variable Z_n does not depend on X_{k+1}, \dots, X_n . Hence,

$$\left| \mathbf{E}\left(Z_n^3 \exp\{it \|\bar{S}_n - a\|^2 + itu Z_n\}\right) \right| \leq \mathbf{E}_k |Z_n|^3 \left| \bar{\mathbf{E}}_k \exp\{it \|\bar{S}_n - a\|^2\} \right|,$$

where as above $\bar{\mathbf{E}}_k(\cdot)$ denotes expectation with respect to X_{k+1}, \dots, X_n with fixed X_1, \dots, X_k . Representing $\bar{S}_n = (\bar{S}_n - \bar{S}_n^k) + \bar{S}_n^k$ we apply Lemma 3 with $Z = \bar{S}_n - \bar{S}_n^k$, $1 \sim n^{2/3}/4$, $m \sim 3n^{2/3}/4$, $A = \frac{3}{4}$. Noting that here $T_1 = |t|$, we obtain

$$(24) \quad \left| \bar{\mathbf{E}}_k \exp\{it \|\bar{S}_n - a\|^2\} \right|$$

$$\leq c_1 \exp(-c_2 n^{2/3}) + c_3 L^{3/2} n^{-1/2} + \prod_{j=1}^7 \left(1 + \frac{c_4 \sigma_j^4 t^2}{n^{2/3}}\right)^{-1/4}.$$

To estimate $\mathbf{E}|Z_n|^3$ we first write

$$(25) \quad \mathbf{E} |Z_n|^3 \leq 32 \varepsilon_n^3 \left(\mathbf{E} |(\bar{S}_n^k, \alpha_k)|^3 + \|a\|^3 \mathbf{E} \|\alpha_k\|^3 \right).$$

Then we represent the scalar product (\bar{S}_n^k, α_k) in the form

$$(26) \quad (\bar{S}_n^k, \alpha_k) = n^{-1/2} \left(\sum_{j=1}^k \bar{g}(X_j), \sum_{1 \leq i_1 < i_2 \leq k} g_2(X_{i_1}, X_{i_2}) \right)$$

$$+ n^{-1/2} k \left(\mathbf{E} \bar{g}_1(X_1), \sum_{1 \leq i_1 < i_2 \leq k} g_2(X_{i_1}, X_{i_2}) \right).$$

Using (8)–(9) we have

$$\mathbf{E} \left| \left(\mathbf{E} \bar{g}_1(X_1), \sum_{1 \leq i_1 < i_2 \leq k} g_2(X_{i_1}, X_{i_2}) \right) \right|^3$$

$$\leq \|\mathbf{E} \bar{g}_1(X_1)\|^3 \mathbf{E} \left\| \sum_{1 \leq i_1 < i_2 \leq k} g_2(X_{i_1}, X_{i_2}) \right\|^3 \leq c_1 n^{-3/2} k^3 \mathbf{E} \|g_2\|^3.$$

To estimate the first term in the right side of (26) we will use the representation

$$\begin{aligned} & \left(\sum_{j=1}^k \bar{g}_1(X_j), \sum_{1 \leq i_1 < i_2 \leq k} g_2(X_{i_1}, X_{i_2}) \right) \\ &= \sum_{j=1}^k \sum_{1 \leq i_1 < i_2 \leq k} (\bar{g}_1(X_j), g_2(X_{i_1}, X_{i_2})) = \sum_{i=1}^5 G_i, \end{aligned}$$

where

$$\begin{aligned} G_1 &= \sum_{j=3}^k \sum_{i_2=2}^{j-1} \sum_{i_1=1}^{i_2-1} (\bar{g}(X_j), g_2(X_{i_1}, X_{i_2})), \\ G_2 &= \sum_{j=2}^k \sum_{i_1=1}^{j-1} (\bar{g}(X_j), g_2(X_{i_1}, X_j)), \\ G_3 &= \sum_{j=2}^{k-1} \sum_{i_1=1}^{j-1} \sum_{i_2=j+1}^k (\bar{g}(X_j), g_2(X_{i_1}, X_{i_2})), \\ G_4 &= \sum_{j=1}^k \sum_{i_2=j+1}^k (\bar{g}(X_j), g_2(X_j, X_{i_2})), \\ G_5 &= \sum_{j=1}^k \sum_{i_2=j+2}^k \sum_{i_1=j+1}^{i_2-1} (\bar{g}(X_j), g_2(X_{i_1}, X_{i_2})). \end{aligned}$$

By (9) we have

$$(27) \quad \begin{aligned} \mathbf{E}|G_i|^3 &\leq c_1 \mathbf{E}\|g_1\|^3 \mathbf{E}\|g_2\|^3 n^{9/2}, \quad i = 1, 3, 5, \\ \mathbf{E}|G_i|^3 &\leq c_2 \mathbf{E}\|g_2\|^3 n^{9/2}, \quad i = 2, 4. \end{aligned}$$

Combining (26) and (27) we obtain

$$\mathbf{E}|Z_n|^3 \leq c_1 (\mathbf{E}\|g_1\|^3 + \|a\|^3) \mathbf{E}\|g_2\|^3 n^{-3/2}.$$

Together with (24) this implies (23).

For $m \sim n$ partition the interval $|t| \leq T = cn^{1/2}L^{-1}$ into the two subsets

$$(28) \quad |t| \leq T(m) = nL^{-1} \left(m \log \frac{m}{L^2} \right)^{-1/2}$$

and

$$(29) \quad T(m) \leq |t| \leq T.$$

LEMMA 5. For t satisfying (28)

$$(30) \quad \begin{aligned} |J_1(t)| &\leq c_1 (1 + \|a\|^3) (\mathbf{E}\|g_1\|^3) (\mathbf{E}\|g_2\|^3)^{1/3} \\ &\quad \times L^3 n^{-1/2} + (|t| + t^2) \prod_{j=1}^7 (1 + c_2 \sigma_j^4 t^2)^{-1/4} \\ &\quad + |t|^3 \prod_{j=1}^9 (1 + c_2 \sigma_j^4 t^2)^{-1/4} \Big) n^{-1/2}, \end{aligned}$$

and for t satisfying (29)

$$(31) \quad |J_1(t)| \leq c_1(1 + \|a\|^3) (\mathbf{E}\|g_1\|^3) (\mathbf{E}\|g_2\|^3)^{1/3} c_9(V) L^2 n^{-1}.$$

Proof. Let us show first that

$$(32) \quad J_1(t) = 4n^{-1/2} \varepsilon_n \binom{k}{2} h_1(t) + 2(k-2) n^{-1/2} \varepsilon_n \binom{k}{2} h_2(t) - 2\varepsilon_n \binom{k}{2} h_3(t),$$

where

$$\begin{aligned} h_1(t) &= \mathbf{E} \left((\bar{g}_1(X_1), g_2(X_1, X_2)) \exp \left\{ it \|\bar{g}_{1n}(X_1) + \bar{g}_{1n}(X_2) + \bar{S}_{n2}\|^2 \right\} \right), \\ h_2(t) &= \mathbf{E} \left((\bar{g}_1(X_1), g_2(X_2, X_3)) \right. \\ &\quad \left. \times \exp \left\{ it \|\bar{g}_{1n}(X_1) + \bar{g}_{1n}(X_2) + \bar{g}_{1n}(X_3) + \bar{S}_{n3}\|^2 \right\} \right), \\ h_3(t) &= \mathbf{E} \left((a, g_2(X_1, X_2)) \exp \left\{ it \|\bar{g}_{1n}(X_1) + \bar{g}_{1n}(X_2) + \bar{S}_{n2}\|^2 \right\} \right), \\ \bar{g}_{1n}(X_j) &= \bar{g}_1(X_j) n^{-1/2}, \quad \bar{S}_{nr} = n^{-1/2} \sum_{j=r+1}^n \bar{g}_1(X_j) - a. \end{aligned}$$

Indeed, according to the definition,

$$\begin{aligned} J_1(t) &= 2n^{-1/2} \varepsilon_n \sum_{i=1}^5 H_i(t) - 2\varepsilon_n H_6, \\ H_1(t) &= \sum_{j=3}^k \sum_{i_2=2}^{j-1} \sum_{i_1=1}^{i_2-1} \mathbf{E} \left((\bar{g}_1(X_j), g_2(X_{i_1}, X_{i_2})) \exp \left\{ it \|\bar{S}_n - a\|^2 \right\} \right) = \binom{k}{3} h_2(t), \\ H_2(t) &= \sum_{j=2}^k \sum_{i_2=1}^{j-1} \mathbf{E} \left((\bar{g}_1(X_j), g_2(X_{i_2}, X_j)) \exp \left\{ it \|\bar{S}_n - a\|^2 \right\} \right) = \binom{k}{2} h_1(t), \\ H_3(t) &= \sum_{j=2}^{k-1} \sum_{i_1=1}^{j-1} \sum_{i_2=j+1}^k \mathbf{E} \left((\bar{g}_1(X_j), g_2(X_{i_1}, X_{i_2})) \exp \left\{ it \|\bar{S}_n - a\|^2 \right\} \right) = \binom{k}{3} h_2(t), \\ H_4(t) &= \sum_{j=1}^{k-1} \sum_{i_1=j+1}^k \mathbf{E} \left((\bar{g}_1(X_j), g_2(X_{i_1}, X_j)) \exp \left\{ it \|\bar{S}_n - a\|^2 \right\} \right) = \binom{k}{2} h_1(t), \\ H_5(t) &= \sum_{j=1}^{k-2} \sum_{i_2=j+2}^k \sum_{i_1=j+1}^{i_2-1} \mathbf{E} \left((\bar{g}_1(X_j), g_2(X_{i_1}, X_{i_2})) \exp \left\{ it \|\bar{S}_n - a\|^2 \right\} \right) \\ &= \binom{k}{3} h_2(t), \\ H_6(t) &= \sum_{i_2=2}^k \sum_{i_1=1}^{i_2-1} \mathbf{E} \left((a, g_2(X_{i_1}, X_{i_2})) \exp \left\{ it \|\bar{S}_n - a\|^2 \right\} \right) = \binom{k}{2} h_3(t). \end{aligned}$$

These representations imply (32). Next we estimate functions $h_i(t)$ in (32). Obviously,

$$\|\bar{g}_{1n}(X_1) + \bar{g}_{1n}(X_2) + \bar{S}_{n2}\|^2 = \|\bar{S}_{n2}\|^2 + W_1 + W_2,$$

where

$$W_1 = \|\bar{g}_{1n}(X_1) + \bar{g}_{1n}(X_2)\|^2, \quad W_2 = 2(\bar{g}_{1n}(X_1) + \bar{g}_{1n}(X_2), \bar{S}_{n2}).$$

By (21) with $r = 1$ and $z = it(W_1 + W_2)$, taking into consideration (2) we have

$$\begin{aligned} h_1(t) &= it \int_0^1 du \mathbf{E} \left((\bar{g}_1(X_1), g_2(X_1, X_2)) W_1 \exp\{itu W_1\} \right. \\ &\quad \left. \times \bar{\mathbf{E}}_2 \exp \left\{ it \|\bar{S}_{n2}\|^2 + itu W_2 \right\} \right) \\ &\quad + it \int_0^1 du \mathbf{E} \left((\bar{g}_1(X_1), g_2(X_1, X_2)) \exp\{itu W_1\} \right. \\ &\quad \left. \times \bar{\mathbf{E}}_2 \left(\exp \left\{ it \|\bar{S}_{n2}\|^2 + itu W_2 \right\} W_2 \right) \right). \end{aligned}$$

The expectation $\bar{\mathbf{E}}_2$ with respect to X_3, \dots, X_n we will estimate using Lemma 3. To estimate $h_1(t)$ in the interval (28) we take in Lemma 3 $l \sim n/2$, $m = n - l$, $A = 1$. Then $T_1 = |t|$ and

$$\begin{aligned} & \left| \bar{\mathbf{E}}_2 \exp \left\{ it \|\bar{S}_{n2}\|^2 + itu W_2 \right\} \right| \\ (33) \quad & \leq c_1 \exp(-c_2 n) + c_3 L^2 n^{-1} + \prod_{j=1}^5 (1 + c_2 \sigma_j^4 t^2)^{-1/4}, \\ & \left| \bar{\mathbf{E}}_2 \left(\exp \left\{ it \|\bar{S}_{n2}\|^2 + itu W_2 \right\} W_2 \right) \right| \leq c_1 (1 + \|a\|) \|\bar{g}_{1n}(X_1) + \bar{g}_{1n}(X_2)\| \\ (34) \quad & \times \exp(-c_2 n) + c_3 L^2 n^{-1} + \prod_{j=1}^5 (1 + c_2 \sigma_j^4 t^2)^{-1/4}. \end{aligned}$$

Moreover,

$$\begin{aligned} & \mathbf{E} \left(\left| (\bar{g}_1(X_1), g_2(X_1, X_2)) \right| \|\bar{g}_{1n}(X_1) + \bar{g}_{1n}(X_2)\| \right) \\ & \leq n^{-1/2} \left(\mathbf{E} \left(\|g_1(X_1)\|^2 \|g_2(X_1, X_2)\| \right) + \mathbf{E} \left(\|g_1(X_1)\| \|g_1(X_2)\| \|g_2(X_1, X_2)\| \right) \right) \\ & \leq 2n^{-1/2} (\mathbf{E} \|g_1\|^3)^{2/3} (\mathbf{E} \|g_2\|^3)^{1/3}, \\ & \mathbf{E} \left(\left| (\bar{g}_1(X_1), g_2(X_1, X_2)) \right| \|\bar{g}_{1n}(X_1) + \bar{g}_{1n}(X_2)\|^2 \right) \\ & \leq 2n^{-1/2} (\mathbf{E} \|g_1\|^3)^{2/3} (\mathbf{E} \|g_2\|^3)^{1/3}. \end{aligned}$$

Thus, by (33) and (34), in the interval (28)

$$\begin{aligned} & |h_1(t)| \leq c_1 (1 + \|a\|) (\mathbf{E} \|g_1\|^3)^{2/3} (\mathbf{E} \|g_2\|^3)^{1/3} \\ (35) \quad & \times \left(Ln^{-1/2} + |t| \prod_{j=1}^7 (1 + c_2 \sigma_j^4 t^2)^{-1/4} \right) n^{-1/2}. \end{aligned}$$

To estimate $h_1(t)$ for t satisfying (29) we take in Lemma 3, $l \sim n/\log n$, $m = n-l$, $A = \frac{5}{4}$. Then $T_1 = T(m)$ and similar reasoning gives for t satisfying (29)

$$\begin{aligned}
 |h_1(t)| &\leq c_1 |t| (1 + \|a\|) (\mathbf{E}\|g_1\|^3)^{2/3} (\mathbf{E}\|g_2\|^3)^{1/3} \\
 &\quad \times \exp\left(-\frac{c_2 n}{\log n}\right) + c_3 L^{5/2} n^{-5/4} (\log n)^{5/4} \\
 &\quad + \prod_{j=1}^5 \left(1 + \frac{c_4 \sigma_j^4 n}{L^2 (\log n) \log(m/L^2)}\right)^{-1/4} n^{-1/2} \\
 (36) \quad &\leq c_5 (1 + \|a\|) (\mathbf{E}\|g_1\|^3)^{2/3} (\mathbf{E}\|g_2\|^3)^{1/3} c_5(V) L^{3/2} n^{-1}.
 \end{aligned}$$

Consider now $h_2(t)$ in (32). Denote

$$\xi = \sum_{j=1}^3 \xi_j, \quad \bar{\xi} = \sum_{j=1}^3 \mathbf{E} \bar{g}_{1n}(X_j), \quad \xi_j = \bar{g}_{1n}(X_j) - \mathbf{E} \bar{g}_{1n}(X_j), \quad j = 1, 2, 3.$$

Obviously,

$$\begin{aligned}
 \|\bar{g}_{1n}(X_1) + \bar{g}_{1n}(X_2) + \bar{g}_{1n}(X_3) + \bar{S}_{n3}\| &= \|\xi + \bar{S}_{n3} + \bar{\xi}\|^2 \\
 &= \|\bar{S}_{n3} + \bar{\xi}\|^2 + W_3 + W_4 + W_5,
 \end{aligned}$$

where

$$W_3 = \sum_{j=1}^3 \|\xi_j\|^2, \quad W_4 = 2 \sum_{j=1}^3 (\xi_j, \bar{S}_{n3} + \bar{\xi}), \quad W_5 = 2 \sum_{1 \leq i < j \leq 3} (\xi_i, \xi_j).$$

Applying (21) with $r = 2$ and $z = it W_5$ we can write

$$(37) \quad h_2(t) = h_{21}(t) + h_{22}(t) + h_{23}(t),$$

where

$$\begin{aligned}
 h_{21}(t) &= \mathbf{E} \left((\bar{g}_1(X_1), g_2(X_2, X_3)) \exp \left\{ it \|\bar{S}_{n3} + \bar{\xi}\|^2 + it(W_3 + W_4) \right\} \right), \\
 h_{22}(t) &= it \mathbf{E} \left((\bar{g}_1(X_1), g_2(X_2, X_3)) W_5 \exp \left\{ it \|\bar{S}_{n3} + \bar{\xi}\|^2 + it(W_3 + W_4) \right\} \right), \\
 h_{23}(t) &= -t^2 \int_0^1 du (1-u) \mathbf{E} \left((\bar{g}_1(X_1), g_2(X_2, X_3)) W_5^2 \right. \\
 &\quad \left. \times \exp \left\{ it \|\bar{S}_{n3} + \bar{\xi}\|^2 + it(W_3 + W_4 + uW_5) \right\} \right).
 \end{aligned}$$

Furthermore, observe that

$$|h_{23}(t)| \leq t^2 \mathbf{E} \left(|(\bar{g}_1(X_1), g_2(X_2, X_3))| W_5^2 \left| \bar{\mathbf{E}}_3 \exp \left\{ it \|\bar{S}_{n3} + \bar{\xi}\|^2 + it W_4 \right\} \right| \right).$$

In the interval (28) by Lemma 3 with $l \sim n/2$, $m = n-l$, $A = \frac{3}{2}$

$$\begin{aligned}
 &\left| \bar{\mathbf{E}}_3 \exp \left\{ it \|\bar{S}_{n3} + \bar{\xi}\|^2 + it W_4 \right\} \right| \\
 (38) \quad &\leq c_1 \exp(-c_2 n) + c_3 L^3 n^{-3/2} + \prod_{j=1}^7 (1 + c_4 \sigma_j^4 t^2)^{-1/4},
 \end{aligned}$$

and for t satisfying (29) Lemma 3 with $l \sim n/\log n$, $m = n - l$, $A = \frac{7}{4}$ gives

$$(39) \quad \begin{aligned} & \left| \bar{\mathbf{E}}_3 \exp \left\{ it \|\bar{S}_{n3} + \bar{\xi}\|^2 + it W_4 \right\} \right| \\ & \leq c_1 \exp \left(-\frac{c_2 n}{\log n} \right) + c_3 L^{7/2} n^{-7/4} (\log n)^{7/4} \\ & \quad + \prod_{j=1}^7 \left(1 + \frac{c_4 \sigma_j^4 n}{L^2 (\log n) \log(m/L^2)} \right)^{-1/4}. \end{aligned}$$

Moreover,

$$(40) \quad \mathbf{E} \left(\left| (\bar{g}_1(X_1), g_2(X_2, X_3)) \right| W_5^2 \right) \leq c \left(\mathbf{E} \|g_1(X_1)\|^3 \right)^{2/3} \left(\mathbf{E} \|g_2\|^3 \right)^{1/3} n^{-3/2}.$$

Combining (38)–(40) we obtain an estimate for $h_{23}(t)$.

By (21) with $r = 1$, $z = it(W_3 + W_4)$ we have

$$(41) \quad h_{22}(t) = F_1(t) + F_2(t),$$

where

$$\begin{aligned} F_1(t) &= it \mathbf{E} \left(\left(\mathbf{E} \bar{g}_1(X_1), g_2(X_2, X_3) \right) W_5 \exp \left\{ it \|\bar{S}_{n3} + \bar{\xi}\|^2 + it (W_3 + W_4) \right\} \right), \\ F_2(t) &= -t^2 \int_0^1 du \mathbf{E} \left(\left(\bar{g}_1(X_1) - \mathbf{E} \bar{g}_1(X_1), g_2(X_2, X_3) \right) W_5 (W_3 + W_4) \right. \\ & \quad \left. \times \exp \left\{ it \|\bar{S}_{n3} + \bar{\xi}\|^2 + itu (W_3 + W_4) \right\} \right). \end{aligned}$$

Here

$$|F_1(t)| \leq |t| \mathbf{E} \left(\left| \left(\mathbf{E} \bar{g}_1(X_1), g_2(X_2, X_3) \right) W_5 \right| \left| \bar{\mathbf{E}}_3 \exp \left\{ it \|\bar{S}_{n3} + \bar{\xi}\|^2 + it W_4 \right\} \right| \right),$$

and

$$\mathbf{E} \left(\left| \left(\mathbf{E} \bar{g}_1(X_1), g_2(X_2, X_3) \right) \right| W_5 \right) \leq 24 \left(\mathbf{E} \|g_2\|^2 \right)^{1/2} n^{-3/2}.$$

Thus for t satisfying (28) by Lemma 3 with $l \sim n/2$, $m = n - l$, $A = 1$

$$(42) \quad \begin{aligned} |F_1(t)| &\leq c_1 |t| \left(\mathbf{E} \|g_2\|^2 \right)^{1/2} \\ &\quad \times c_1 \exp(-c_2 n) + c_3 L^2 n^{-1} + \prod_{j=1}^5 \left(1 + c_4 \sigma_j^4 t^2 \right)^{-1/4} \Big) n^{-3/2}, \end{aligned}$$

and for t satisfying (29) by Lemma with $l \sim n/\log n$, $m = n - l$ and $A = \frac{5}{4}$,

$$(43) \quad \begin{aligned} |F_1(t)| &\leq c_1 |t| \left(\mathbf{E} \|g_2\|^2 \right)^{1/2} c_1 \left(\exp \left(-\frac{c_2 n}{\log n} \right) + c_3 L^{5/2} n^{-5/4} (\log n)^{5/4} \right. \\ & \quad \left. + \prod_{j=1}^5 \left(1 + \frac{c_4 \sigma_j^4 n}{L^2 (\log n) \log(m/L^2)} \right)^{-1/4} \right) n^{-3/2}. \end{aligned}$$

To estimate $f_2(t)$ in (41) we first observe, that

$$|F_2(t)| \leq t^2 \int_0^1 du \mathbf{E} \left(\left| (\bar{g}_1(X_1) - \mathbf{E}\bar{g}_1(X_1), g_2(X_2, X_3)) W_5 \right| \right. \\ \left. \times \left| \bar{\mathbf{E}}_3(W_3 + W_4) \exp \left\{ it \|\bar{S}_{n3} + \bar{\xi}\|^2 + itu W_4 \right\} \right| \right),$$

and

$$\mathbf{E} \left(\left| (\bar{g}_1(X_1) - \mathbf{E}\bar{g}_1(X_1), g_2(X_2, X_3)) W_5 \right| \left[W_3 + \sum_{j=1}^3 \|\xi_j\| (1 + \|a\|) \right] \right) \\ \leq c \mathbf{E}\|g_1\|^3 (\mathbf{E}\|g_2\|^3)^{1/3} (1 + \|a\|) n^{-3/2}.$$

Applying now Lemma 3 for t satisfying (28) with $l \sim n \log n$, $m = n - l$ and $A = \frac{3}{2}$, we have

$$(44) \quad |F_2(t)| \leq c_1 t^2 (1 + \|a\|) \mathbf{E}\|g_1\|^3 (\mathbf{E}\|g_2\|^3)^{1/3} \\ \times \exp(-c_2 n) c_3 L^3 n^{-3/2} + \prod_{j=1}^7 (1 + c_4 \sigma_j^4 t^2)^{-1/4} \Big) n^{-3/2},$$

and for t satisfy (29), using Lemma 3 with $l \sim n \log n$, $m = n - l$ and $A = \frac{7}{4}$, we have

$$(45) \quad |F_2(t)| \leq c_1 t^2 (1 + \|a\|) \mathbf{E}\|g_1\|^3 (\mathbf{E}\|g_2\|^3)^{1/3} \\ \times \exp \left(-\frac{c_2 n}{\log n} \right) + c_3 L^{7/2} n^{-7/4} (\log n)^{7/4} \\ + \prod_{j=1}^7 \left(1 + \frac{c_4 \sigma_j^4 n}{L^2 (\log n) \log(m/L^2)} \right)^{-1/4} \Big) n^{-3/2}.$$

In (37) it remains only to estimate $h_{21}(t)$. We use the representation

$$(46) \quad h_{21}(t) = F_3(t) + F_4(t),$$

where

$$F_3(t) = \mathbf{E} \left((\mathbf{E}\bar{g}_1(X_1), g_2(X_2, X_3)) \exp \left\{ it \|\bar{S}_{n3} + \bar{\xi}\|^2 + it (W_3 + W_4) \right\} \right), \\ F_4(t) = \mathbf{E} \left((\bar{g}_1(X_1) - \mathbf{E}\bar{g}_1(X_1), g_2(X_2, X_3)) \right. \\ \left. \times W_5 \exp \left\{ it \|\bar{S}_{n3} + \bar{\xi}\|^2 + it (W_3 + W_4) \right\} \right).$$

In view of (2) we can also write

$$(47) \quad F_3(t) = \mathbf{E} \left((\mathbf{E}\bar{g}_1(X_1), g_2(X_2, X_3)) \exp \left\{ it \|\bar{S}_{n3} + \bar{\xi}\|^2 + it \|\xi_1\|^2 + 2it (\xi_1, \bar{S}_{n3} + \bar{\xi}) \right\} \right. \\ \left. \times \prod_{j=2}^3 \left(\exp \left\{ it \|\xi_j\|^2 + 2it (\xi_j, \bar{S}_{n3} + \bar{\xi}) \right\} - 1 \right) \right),$$

$$\begin{aligned}
F_4(t) &= \mathbf{E} \left(\bar{g}_1(X_1) - \mathbf{E}\bar{g}_1(X_1), g_2(X_2, X_3) \right) \exp \left\{ it \|\bar{S}_{n3} + \bar{\xi}\|^2 \right\} \\
(48) \quad &\times \prod_{j=1}^3 \left(\exp \left\{ it \|\xi_j\|^2 + 2it(\xi_j, \bar{S}_{n3} + \bar{\xi}) \right\} - 1 \right).
\end{aligned}$$

Furthermore, applying (21) with $r = 1$ in (47) and (48), we have

$$\begin{aligned}
F_3(t) &= -t^2 \mathbf{E} \left(\mathbf{E}\bar{g}_1(X_1), g_2(X_2, X_3) \right) \\
&\times \exp \left\{ it \|\bar{S}_{n3} + \bar{\xi}\|^2 + it \|\xi_1\|^2 + 2it(\xi_1, \bar{S}_{n3} + \bar{\xi}) \right\} \\
&\times \prod_{j=2}^3 \left(\|\xi_j\|^2 + 2(\xi_j, \bar{S}_{n3} + \bar{\xi}) \right) \\
(49) \quad &\times \int_0^1 du \exp \left\{ itu \|\xi_j\|^2 + 2itu(\xi_j, \bar{S}_{n3} + \bar{\xi}) \right\},
\end{aligned}$$

$$\begin{aligned}
F_4(t) &= (it)^2 \mathbf{E} \left(\bar{g}_1(X_1) - \mathbf{E}\bar{g}_1(X_1), g_2(X_2, X_3) \right) \exp \left\{ it \|\bar{S}_{n3} + \bar{\xi}\|^2 \right\} \\
&\times \prod_{j=1}^3 \left(\|\xi_j\|^2 + 2(\xi_j, \bar{S}_{n3} + \bar{\xi}) \right) \\
(50) \quad &\times \int_0^1 du \exp \left\{ itu \|\xi_j\|^2 + 2itu(\xi_j, \bar{S}_{n3} + \bar{\xi}) \right\}.
\end{aligned}$$

Denote

$$b_j = \|\xi_j\|^2 + 2(\xi_j, \bar{\xi}) - 2(\xi_j, a).$$

Then

$$\begin{aligned}
&\prod_{j=2}^3 \left(\|\xi_j\|^2 + 2(\xi_j, \bar{S}_{n3} + \bar{\xi}) \right) = \prod_{j=2}^3 \left(b_j + 2(\xi_j, S_{n3} + a) \right) \\
(51) \quad &= b_2 b_3 + 2b_2(\xi_3, \bar{S}_{n3} + a) + 2b_3(\xi_2, \bar{S}_{n3} + a) + 4(\xi_2, \bar{S}_{n3} + a)(\xi_3, \bar{S}_{n3} + a),
\end{aligned}$$

where according to the notation introduced above

$$\bar{S}_{n3} + a = n^{-1/2}(\bar{g}_1(X_4) + \dots + \bar{g}_1(X_n)).$$

Note that

$$\begin{aligned}
(52) \quad &\|\xi_j\| \leq 2, \quad \|\xi_j\| \leq n^{-1/2} \left(1 + \|g_1(X_j)\| \right), \\
&|b_j| \leq 4(1 + \|a\|) \left(1 + \|g_1(X_j)\| \right) n^{-1/2}.
\end{aligned}$$

Substituting now (51) in (49) we obtain four expectations. Then we apply Lemma 3 to each of them, taking into account (52) and the estimate $\|\mathbf{E}\bar{g}_1(X_1)\| \leq n^{-1/2}$. For t satisfying (28) we take in Lemma 3 $l \sim n/2$, $m = n - l$, $A = \frac{3}{2}$ and obtain

$$|F_3(t)| \leq c_1 t^2 (1 + \|a\|)^2 (\mathbf{E}\|g_2\|^2)^{1/2}$$

$$(53) \quad \times \exp(-c_2 n) + c_3 L^3 n^{-3/2} + \prod_{j=1}^7 (1 + c_4 \sigma_j^4 t^2)^{-1/4} \Big) n^{-3/2},$$

and for t satisfying (29) we take $l \sim n/\log n$, $m = n - l$ and $A = \frac{7}{4}$ and obtain

$$(54) \quad \begin{aligned} |F_3(t)| &\leq c_1 t^2 (1 + \|a\|)^2 (\mathbf{E}\|g_2\|^2)^{1/2} \\ &\times \exp\left(-\frac{c_2 n}{\log n}\right) + c_3 L^{7/2} n^{-7/4} (\log n)^{7/4} \\ &+ \prod_{j=1}^7 \left(1 + \frac{c_4 \sigma_j^4 n}{L^2 (\log n) \log(n/L^2)}\right)^{-1/4} \Big) n^{-3/2}. \end{aligned}$$

To estimate $f_4(t)$ in (50) we use the following formula which is analogous to (51):

$$\begin{aligned} \prod_{j=1}^3 &\left(\|\xi_j\|^2 + 2(\xi_j, \bar{S}_{n_3} + \bar{\xi})\right) = b_1 b_2 b_3 + 2b_1 b_2 (\xi_3, \bar{S}_{n_3} + a) \\ &+ 2b_1 b_3 (\xi_2, \bar{S}_{n_3} + a) + 2b_2 b_3 (\xi_1, \bar{S}_{n_3} + a) + 4b_1 (\xi_2, \bar{S}_{n_3} + a) (\xi_3, \bar{S}_{n_3} + a) \\ &+ 4b_2 (\xi_1, \bar{S}_{n_3} + a) (\xi_3, \bar{S}_{n_3} + a) + 4b_3 (\xi_1, \bar{S}_{n_3} + a) (\xi_2, \bar{S}_{n_3} + a) \\ &+ 8(\xi_1, \bar{S}_{n_3} + a) (\xi_2, \bar{S}_{n_3} + a) (\xi_3, \bar{S}_{n_3} + a). \end{aligned}$$

Substituting it in (50) we obtain eight similar expectations. To estimate these expectations we again apply Lemma 3. For t satisfying (28) we take in it $l \sim n/2$, $m = n - l$, $A = 2$ and obtain,

$$(55) \quad \begin{aligned} |F_4(t)| &\leq c_1 |t|^3 (1 + \|a\|)^3 (\mathbf{E}\|g_2\|^2)^{1/2} \\ &\times \exp(-c_2 n) + c_3 L^4 n^{-2} + \prod_{j=1}^9 (1 + c_4 \sigma_j^4 t^2)^{-1/4} \Big) n^{-3/2}, \end{aligned}$$

and for t satisfying (29) we take $l \sim n/\log n$, $m = n - l$ and $A = \frac{9}{4}$ and obtain

$$(56) \quad \begin{aligned} |F_4(t)| &\leq c_1 |t|^3 (1 + \|a\|)^3 (\mathbf{E}\|g_2\|^2)^{1/2} \\ &\times \exp\left(-\frac{c_2 n}{\log n}\right) + c_3 L^{9/2} n^{-9/4} (\log n)^{9/4} \\ &+ \prod_{j=1}^9 \left(1 + \frac{c_4 \sigma_j^4 n}{L^2 (\log n) \log(m/L^2)}\right)^{-1/4} \Big) n^{-3/2}. \end{aligned}$$

Estimates (37)–(38), (41)–(42), (53), and (55) together for t satisfying (28) imply

$$(57) \quad \begin{aligned} |h_2(t)| &\leq c_1 (1 + \|a\|)^3 (\mathbf{E}\|g_1\|^3) (\mathbf{E}\|g_2\|^3)^{1/3} \\ &\times Ln^{-1/2} + (|t| + t^2) \prod_{j=1}^7 (1 + c_4 \sigma_j^4 t^2)^{-1/4} \\ &+ |t|^3 \prod_{j=1}^9 (1 + c_4 \sigma_j^4 t^2)^{-1/4} \Big) n^{-3/2}, \end{aligned}$$

and (37), (39), (41), (43), (54), and (56) for t satisfying (29) imply

$$(58) \quad |h_2(t)| \leq c_1(1 + \|a\|)^3 (\mathbf{E}\|g_1\|^3) (\mathbf{E}\|g_2\|^3)^{1/3} c_9(V) L^{3/2} n^{-2}.$$

It remains to estimate $h_3(t)$ in (32). Applying (21) with $r = 1$ and $z = it(W_1 + W_2)$, $z = itu(W_1 + W_2)$ twice, we have

$$\begin{aligned} h_3(t) = & it \int_0^1 du \mathbf{E} \left((a, g_2(X_1, X_2)) W_1 \exp\{itu W_1\} \bar{\mathbf{E}}_2 \exp \left\{ it \|\bar{S}_{n2}\|^2 + itu W_2 \right\} \right) \\ & - t^2 \int_0^1 dv \int_0^1 du u \mathbf{E} \left((a, g_2(X_1, X_2)) \exp\{ituv W_1\} \right. \\ & \left. \times \bar{\mathbf{E}}_2 \left(\exp \left\{ it \|\bar{S}_{n2}\|^2 + ituv W_2 \right\} (W_1 + W_2) \right) \right). \end{aligned}$$

Arguing in the same way as in proving (35) and (36) we obtain for t satisfying (28)

$$(59) \quad |h_3(t)| \leq c_1(1 + \|a\|)^3 (\mathbf{E}\|g_1\|^3)^{2/3} (\mathbf{E}\|g_2\|^3)^{1/3} \times \left[Ln^{-1/2} + (|t| + t^2) \prod_{j=1}^7 (1 + c_2 \sigma_j^4 t^2)^{-1/4} \right] n^{-1},$$

and for t satisfying (29)

$$(60) \quad |h_3(t)| \leq c_1(1 + \|a\|^3) (\mathbf{E}\|g_1\|^3)^{2/3} (\mathbf{E}\|g_2\|^3)^{1/3} c_7(V) L^{3/2} n^{-3/2}.$$

Now (32), (35), (57), (59) imply (30) and (32), (36), (58), (60) imply (31).

LEMMA 6. *If t belongs to the interval (28), then*

$$(61) \quad |J_2(t)| \leq c_1(1 + \|a\|^2) (\mathbf{E}\|g_1\|^3)^{2/3} \times L^2 n^{-1/2} + n^{1/6} \prod_{j=1}^9 \left(1 + \frac{c_2 \sigma_j^4 t^2}{n^{1/3}} \right)^{-1/6} n^{-1},$$

and if t satisfies (29)

$$(62) \quad |J_2(t)| \leq c_1(1 + \|a\|^2) (\mathbf{E}\|g_1\|^3)^{2/3} c_9(V) L^3 n^{-3/2}.$$

Proof. Noting that

$$J_2(t) = \mathbf{E} \left(Z_n^2 \bar{\mathbf{E}} \exp \{ it \|\bar{S}_n - a\|^2 \} \right)$$

we have

$$(63) \quad |J_2(t)| \leq 4\varepsilon_n^2 \mathbf{E} \left(\|\bar{S}_n^k\|^2 \|\alpha_k\|^2 |\bar{\mathbf{E}}_k \exp \{ it \|\bar{S}_n - a\|^2 \}| \right) + 4\varepsilon_n^2 \|a\|^2 \mathbf{E} \left(\|\alpha_k\|^2 |\bar{\mathbf{E}}_k \exp \{ it \|\bar{S}_n - a\|^2 \}| \right).$$

By Hölder's inequality

$$(64) \quad \begin{aligned} & \mathbf{E} \left(\|\bar{S}_n^k\|^2 \|\alpha_k\|^2 |\bar{\mathbf{E}}_k \exp \{ it \|\bar{S}_n - a\|^2 \}| \right) \\ & \leq (\mathbf{E}\|\alpha_k\|^3)^{2/3} \left(\mathbf{E} \left(\|\bar{S}_n^k\|^6 |\bar{\mathbf{E}}_k \exp \{ it \|\bar{S}_n - a\|^2 \}|^3 \right) \right)^{1/3}. \end{aligned}$$

Since for any $x > 0$ and $p > 0$

$$\|\overline{S}_n^k\|^6 \leq x^6 + x^{-p} \|\overline{S}_n^k\|^{p+6},$$

we have in (64)

$$\begin{aligned} & \mathbf{E} \left(\|\overline{S}_n^k\|^6 \left| \overline{\mathbf{E}}_k \exp \{it \|\overline{S}_n - a\|^2\} \right|^3 \right) \\ (65) \quad & \leq x^6 \mathbf{E} \left| \overline{\mathbf{E}}_k \exp \{it \|\overline{S}_n - a\|^2\} \right|^2 + x^{-p} \mathbf{E} \|\overline{S}_n^k\|^{p+6}. \end{aligned}$$

To estimate the second expectation in (63) we also use Hölder's inequality and obtain

$$\begin{aligned} & \mathbf{E} \left(\|\alpha_k\|^2 \left| \overline{\mathbf{E}}_k \exp \{it \|\overline{S}_n - a\|^2\} \right| \right) \leq \mathbf{E} (\|\alpha_k\|^3)^{2/3} \left(\mathbf{E} \left| \overline{\mathbf{E}}_k \exp \{it \|\overline{S}_n - a\|^2\} \right|^3 \right)^{1/3} \\ (66) \quad & \leq \mathbf{E} (\|\alpha_k\|^3)^{2/3} \left(\mathbf{E} \left| \overline{\mathbf{E}}_k \exp \{it \|\overline{S}_n - a\|^2\} \right|^2 \right)^{1/3}. \end{aligned}$$

Substituting (64) in (63) and using the elementary inequality $(\alpha + \beta)^{1/3} \leq \alpha^{1/3} + \beta^{1/3}$ we get

$$\begin{aligned} & |J_2(t)| \leq 4\epsilon_n^2 \mathbf{E} (\|\alpha_k\|^3)^{2/3} \\ (67) \quad & \times \left[x^{-p/3} (\mathbf{E} \|\overline{S}_n^k\|^{p+6})^{1/3} + (x^2 + \|a\|^2) \left(\mathbf{E} \left| \overline{\mathbf{E}}_k \exp \{it \|\overline{S}_n - a\|^2\} \right|^2 \right)^{1/3} \right]. \end{aligned}$$

Applying the same argument as in the proof of Lemma 10 in [3] with $l \sim n^{2/3}$, $m = n - l$, $A = \frac{3}{2}$ for $|t| \leq cn^{1/2}L^{-1}$, we have

$$\mathbf{E} \left| \overline{\mathbf{E}}_k \exp \{it \|\overline{S}_n - a\|^2\} \right|^2 \leq c_1 L^3 n^{-1} + \prod_{j=1}^9 \left(1 + \frac{c_2 \sigma_j^4 T_1^2}{n^{1/3}} \right)^{-1/2},$$

where T_1 is defined in Lemma 3. Furthermore, taking in (67) $p = 36$, $x = n^{1/12}$ and using (8) and (9) we obtain the estimate

$$\begin{aligned} & |J_2(t)| = c_1 (1 + \|a\|^2) (\mathbf{E} \|g_2\|^3)^{2/3} \\ & \times L^2 n^{-1/2} + n^{1/6} \prod_{j=1}^9 \left(1 + \frac{c_2 \sigma_j^4 T_1^2}{n^{1/3}} \right)^{-1/6} n^{-1}, \end{aligned}$$

which implies (61) and (62).

LEMMA 7. I_2 in (19) satisfies the inequality

$$(68) \quad I_2 \leq c_1 (1 + \|a\|^3) \beta^2 \sigma_9^{-1} c_9(V) n^{-1/2}.$$

Proof. (22) implies

$$(69) \quad |\varphi_n(t) - \psi_n(t)| |t|^{-1} \leq |J_1(t)| + |t| |J_2(t)| + t^2 |J_3(t)|.$$

By Lemma 4

$$(70) \quad \int_{-T}^T t^2 |J_3(t)| dt \leq c_1 (1 + \|a\|^3) \mathbf{E} \|g_1\|^3 \mathbf{E} \|g_2\|^3 c_7(V) n^{-1/2}.$$

By Lemma 5

$$(71) \quad \int_{-T}^T |J_1(t)| dt \leq c_1(1 + \|a\|^3) (\mathbf{E}\|g_1\|^3) (\mathbf{E}\|g_2\|^3)^{1/3} c_9(V) L^{1/2} n^{-1/2}.$$

By Lemma 6

$$(72) \quad \int_{-T}^T |t| |J_2(t)| dt \leq c_1(1 + \|a\|^2) (\mathbf{E}\|g_1\|^3)^{2/3} c_9(V) Ln^{-1/2}.$$

Noting now that the moments in (70)–(72) satisfy (4) and integrating both sides in (69), we obtain (68).

From (14), (19)–(20), and (68) it follows that

$$(73) \quad \Delta_{n2}(a) \leq c_1(1 + \|a\|^3) \beta^2 \sigma_9^{-1} c_9(V) n^{-1/2}.$$

Estimates (5), (6), and (73) imply (3) and the theorem is proved.

REFERENCES

- [1] YU. V. BOROVSKIKH, *Theory of U-statistics in Hilbert space*, Preprint No 86.78, Institute of Mathematics, Ukraine Academy of Science, Kiev, 1986 (in Russian).
- [2] YU. V. BOROVSKIKH AND V. S. KOROLYUK, *UH-statistics*, Soviet Math. Dokl., 40 (1990), pp. 432–435.
- [3] B. A. ZALESSKII, V. V. SAZONOV, AND V. V. UL'YANOV, *Normal approximation in Hilbert space*, I, II, Theory Probab. Appl., 33 (1988), pp. 207–227; pp. 473–487.
- [4] V. S. KOROLYUK AND YU. V. BOROVSKIKH, *Approximation of nondegenerate U-statistics*, Theory Probab. Appl., 30 (1986), pp. 439–450.
- [5] V. S. KOROLYUK AND YU. V. BOROVSKIKH, *Theory of U-statistics*, Kluwer, Dordrecht, 1994.
- [6] V. PAULASKAS AND A. RAČKAUSKAS, *Approximation Theory in the Central Limit Theorem. Exact Results in Banach Spaces*, Kluwer, Dordrecht, 1989.
- [7] V. V. PETROV, *Sums of Independent Random Variables*, Springer-Verlag, Berlin, 1975.
- [8] I. F. PINELIS, *Estimates of moments of infinite dimensional martingales*, Math. Notes, 27 (1980), pp. 459–462.
- [9] B. A. ZALESSKII, V. V. SAZONOV, AND V. V. UL'YANOV, *A precise estimate of the rate of convergence in the central limit theorem in Hilbert space*, Math. USSR Sbornik, 68 (1991), pp. 453–482.
- [10] V. BENTKUS, F. GÖTZE, AND R. ZITIKIS, *Lower estimates of the convergence rate for U-statistics*, Ann. Probab., 22 (1994), No. 4.
- [11] W. S. GRAMS AND R. J. SERFLING, *Convergence rate for U-statistics.*, Ann. Statist., 1 (1973), pp. 153–160.
- [12] V. S. KOROLYUK AND YU. V. BOROVSKIKH, *Rate of convergence in the central limit theorem for UH-statistics*, Theory Probab. Math. Statist., 43 (1991), pp. 79–85.
- [13] A. J. LEE, *U-statistics; Theory and Practice*, Marcel Dekker, New York, 1990.
- [14] M. L. PURI AND V. V. SAZONOV, *On Hilbert space valued U-statistics*, Theory Probab. Appl., 36 (1991), pp. 604–605.
- [15] V. V. SAZONOV, *Normal Approximation — Some Recent Advances*, Lecture Notes in Math., Vol. 879, Springer-Verlag, Berlin, 1981.
- [16] V. V. SAZONOV, V. V. UL'YANOV, AND B. A. ZALESSKII, *Asymptotically precise estimate of accuracy of Gaussian approximation in Hilbert space*, J. Multivariate Anal., 28 (1989), pp. 304–330.
- [17] V. V. SAZONOV AND V. V. UL'YANOV, *Speed of convergence in the central limit theorem in Hilbert space under weakened moment conditions*, in Proc. 5th Vilnius Conference on Probability Theory and Mathematical Statistics, 1989, Mokslas/VSP, Vilnius/Utrecht, 1990, pp. 394–410.